#### STAT 315: Joint PDF of continuous random variables

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## Bivariate or joint random variables

Suppose we have random experiment and make TWO measurements  $Y_1$  and  $Y_2$  or more....

#### **Examples:**

- We measure the height and weight of some individual in a population.
- We have an exponential random variable whose parameter  $\beta$  is itself random and obeys a certain distribution.
- We throw a dart at a random position  $(Y_1, Y_2)$  on a circular target.
- ...
- Sampling: We repeat an experiment n times and record the results  $Y_1, Y_2, \dots, Y_n$  of the experiments (the most important example!)

We need to describe the probability distribution of  $Y_1$  and  $Y_2$  together! This is called the joint (or bivariate) PDF  $f(y_1, y_2)$  (continuous).

#### Joint PDF of continuous

#### Joint (or bivariate) PDF for continuous random variables

The joint continuous RV  $(Y_1, Y_2)$  have joint PDF  $f(y_1, y_2)$  if

$$P(a_1 \le Y_1 \le b_1, a_2 \le Y_2 \le b_2) = \int_{a_1}^{b_1} \int_{a_2}^{b_2} f(y_1, y_2) dy_1, dy_2$$
with  $0 \le f(y_1, y_2)$  and  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(y_1, y_2) dy_1 dy_2 = 1$ 

**Example 1:** The random variables (X, Y) have the joint PDF

$$f(x,y) = \begin{cases} 2e^{-2x}e^{-y} & \text{if } x \ge 0, y \ge 0\\ 0 & \text{else} \end{cases}$$

**Example 2:** A gas station adds some gas to its tank every Monday morning. We write  $Y_1$  (in [0,1]) for the proportion of the tank being filled and  $Y_2$  for the proportion of the tank which is sold to customers during the subsequent. Note that we must have  $Y_2 \leq Y_1$ . We propose the model

$$f(y_1, y_2) = \begin{cases} 3y_1 & \text{if } 0 \le y_2 \le y_1 \le 1\\ 0 & \text{else} \end{cases}$$

**Example 3:** Two friends, independently of each other, arrive at a random time between 12pm and 1pm at the blue wall. We can describe the time of their arrival by two random variable  $Y_1$ ,  $Y_2$  each with a uniform RV on [0,1] (measured in hours). The independence assumption leads to the model

$$f(y_1, y_2) = \begin{cases} 1 & \text{if } 0 \le y_1 \le 1, 0 \le y_2 \le 1 \\ 0 & \text{else} \end{cases}$$

**Example 4:** A (pretty bad) player throws a dart at a random point on a circular target of radius R. This can be described by a uniform distribution on a disk of radius R that is by the joint random variables  $(X_1, X_2)$  with pdf

$$f(x_1, x_2) = \begin{cases} \frac{1}{\pi R^2} & \text{if } x_1^2 + x_2^2 \le R^2 \\ 0 & \text{else} \end{cases}$$

**Example 5:** In Bayesian statistics context one uses random variables whose parameters are themselves random variables. For example consider the joint PDF

$$f(x,y) = \begin{cases} ye^{-yx}e^{-y} & \text{if } x \ge 0, y \ge 0\\ 0 & \text{else} \end{cases}$$

As we will see this describe a an exponential random variable whose scale parameters (i.e. with pdf  $\lambda e^{-\lambda x}$ ) has a exponential distribution with parameter 1.

# Marginal and conditional PDF

### Marginal PDF of continuous random variables

If the joint continuous RV  $(Y_1, Y_2)$  has PDF  $f(y_1, y_2)$  then the marginal PDFs of  $Y_1$  and  $Y_2$  are given by

$$f(y_1) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_2$$
  $f(y_2) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_1$ 

#### Conditional PDF of continuous random variables

If the joint continuous RV  $(Y_1, Y_2)$  has PDF  $p(y_1, y_2)$  then the conditional PDFs of  $Y_1$  given  $Y_2 = y_2$  is given by

$$f(y_1|y_2) = \frac{f(y_1, y_2)}{f(y_2)}$$

if 
$$f(y_2) > 0$$

### Independence

Recall that the events A and B are independent if

$$P(A|B) = P(A)$$
 or  $P(B|A) = P(B)$  or  $P(A \cap B) = P(A)P(B)$ 

### Independence of continuous random variables

The continuous random variables  $Y_1$  and  $Y_2$  are independent if

$$f(y_1|y_2) = f(y_1)$$
 or  $f(y_2|y_1) = f(y_2)$  or  $f(y_1, y_2) = f(y_1)f(y_2)$ 

#### Criterion for independence

The random variables  $Y_1$  and  $Y_2$  are independent if and only if

$$f(y_1, y_2) = g(y_1)h(y_2) - \infty < y_1, y_2 < \infty$$

for some function g(x) and h(y)

# Expected value of function of joint random variables

#### Expected value

For joint random variables  $Y_1$  and  $Y_2$  and a function  $g(Y_1, Y_2)$  we have

$$E[g(Y_1, Y_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(y_1, y_2) f(y_1, y_2) dy_1 dy_2 \quad \text{continuous RV}$$

with properties

#### Linearity of expectation

- $\bullet$  E[c] = c
- $E[c g(Y_1, Y_2)] = c E[g(Y_1, Y_2)]$
- $E[g(Y_1, Y_2) + h(Y_1, Y_2)] = E[g(Y_1, Y_2)] + E[h(Y_1, Y_2)]$

## Independence and products

### Independence and products

If  $Y_1$  and  $Y_2$  are independent then for any functions  $g(Y_1)$  and  $h(Y_2)$ 

$$E[g(Y_1)h(Y_2)] = E[g(Y_1)]E[h(Y_2)]$$

For example independence implies that have

$$E[Y_1Y_2] = E[Y_1]E[Y_2]$$

**Proof:** 

#### Covariance

### Covariance of $Y_1$ and $Y_2$

If  $Y_1$  and  $Y_2$  are random variables with means  $\mu_1 = E[Y_1]$  and  $\mu_2 = E[Y_2]$  then the covariance of  $Y_1$  and  $Y_2$  is

$$Cov(Y_1, Y_2) = E[(Y_1 - \mu_1)(Y_2 - \mu_2)]$$

and the correlation coefficient  $\rho$  is

$$\rho = \rho(Y_1, Y_2) = \frac{\operatorname{Cov}(Y_1, Y_2)}{\sigma_1 \sigma_2}$$

We say that  $Y_1$  and  $Y_2$  are

- positively correlated if  $Cov(Y_1, Y_2) > 0$
- negatively correlated if  $Cov(Y_1, Y_2) < 0$
- uncorrelated if  $Cov(Y_1, Y_2) = 0$

### Properties of covariance

We have the formula

$$Cov(Y_1, Y_2) = E[Y_1Y_2] - E[Y_1]E[Y_2]$$

- ②  $Cov(Y_1, Y_1) = V(Y_1)$  and so  $\rho(Y_1, Y_1) = 1$
- We have Cauchy-Schwartz inequality

$$|E[Z_1Z_2]| \le \sqrt{E[Z_1^2]E[Z_2^2]}$$

and as a consequence the correlation coefficient satisfies

$$-1 \le \rho \le 1$$

• If  $Y_1$  and  $Y_2$  are independent then  $Cov(Y_1, Y_2) = 0$  and so  $Y_1$  and  $Y_2$  are uncorrelated. But the converse is not always true

#### Linear combinations of random variables

For random variables  $Y_1$ ,  $Y_2$  and  $Z_1$ ,  $Z_2$  and constants  $a_1$ ,  $a_2$  and  $b_1$ ,  $b_2$ .

Expected Value

$$E[a_1Y_1 + a_2Y_2] = a_1E[Y_1] + a_2E[Y_2]$$

Variance

$$V(a_1Y_1 + a_2Y_2) = a_1^2V(Y_1) + a_2^2V(Y_2) + 2a_1a_2\text{Cov}(Y_1, Y_2)$$

Covariance

$$Cov(a_1Y_1 + a_2Y_2, b_1Z_1 + b_2Z_2) = a_1b_1Cov(Y_1, Z_1) + a_1b_2Cov(Y_1, Z_2) + a_2b_1Cov(Y_2, Z_1) + a_2b_2Cov(Y_2, Z_2)$$

### Example

- Suppose  $V[Y_1] = 2$  and  $V[Y_2] = 8$ . What are the possible values for the covariance  $Cov(Y_1, Y_2)$ ?
- What is Cov(1, *Y*)?
- What is Cov(1 + Y, 3 Y)?
- Suppose  $\rho(Y_1, Y_2) = .2$ . What is  $\rho(1 + Y_1, 3 2Y_2)$
- Suppose X and Y are independent with variance  $\sigma_X^2$  and  $\sigma_Y^2$ . What is V[X-Y]
- In Example 2.  $U = Y_1 Y_2$  is the proportion of unsold gas. Compute the mean and the variance of U.

## Illustration of covariances: diversifying your investment

- Think of  $X_1$  and  $X_2$  has the return on investment two investment strategies like investing in stocks, or in bonds, or in crypto, and so on.... If you invest 1 unit in startegy i then you will make a profit  $X_i$ .
- For simplicity think of all these strategies are equally good. That is we have  $E[X_1] = E[X_2] = \mu$ .
- ullet Diversification strategy, pick  $0 \le \alpha < 1$  and follow the strategy

$$X_{\alpha} = \alpha X_1 + (1 - \alpha) X_2.$$

We have

$$E[X_{\alpha}] = E[\alpha X_1 + (1 - \alpha)X_2] = \alpha \mu + (1 - \alpha)\mu = \mu$$

so they all give the same return, no matter what  $\alpha$  is.

• To chose the less risky investment we need to analyze the variance.

• If  $X_1$  and  $X_2$  are independent we have

$$V[\alpha X_1 + (1 - \alpha)X_2] = \alpha \sigma_1^2 + (1 - \alpha)^2 \sigma_2^2$$

Differentiating to find the optimal  $\alpha$  we have

$$2\alpha\sigma_1^2 - 2(1-\alpha)\sigma_2^2 = 0 \implies \alpha^* = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2}, 1 - \alpha^* = \frac{\sigma_1^2}{\sigma_1^2 + \sigma_2^2}$$

For example if  $\sigma_1^2=1$  and  $\sigma_2^2=4$  then  $lpha^*=\frac{4}{5}$  and

$$V[X_{\alpha^*}] = (\frac{4}{5})^2 \times 1 + (\frac{1}{5})^2 \times 4 = \frac{4}{5} < 1$$

so the optimal variance is smaller than the smallest variance! In general

$$V[X_{\alpha^*}] = \frac{\sigma_4^2 \sigma_1^2 + \sigma_2^2 \sigma_1^4}{(\sigma_1^2 + \sigma_2^2)^2} = \frac{\sigma_2^2 \sigma_1^2}{(\sigma_1^2 + \sigma_2^2)} < \frac{\sigma_1^2}{(\sigma_1^2 + \sigma_2^2)}$$
(or  $< \sigma_2^2$ )

• If  $X_1$  and  $X_2$  are negatively correlated then it is even better. We have

$$Cov(X_1, X_2) = \rho \sigma_1 \sigma_2 < 0$$

and so

$$V[X_{\alpha}] = \alpha^{2}\sigma_{1}^{2} + (1-\alpha)^{2}\sigma_{2}^{2} + 2\alpha(1-\alpha)\rho\sigma_{1}\sigma_{2} < \alpha^{2}\sigma^{1} + (1-\alpha)^{2}\sigma_{2}^{2}$$

ullet For example if  $\sigma_1^2=1$ ,  $\sigma_2^2=4$ , and  $ho=-\frac{1}{2}$  then we have

$$V[X_{\alpha}] = \alpha^2 + 4(1 - \alpha)^2 - 2\alpha(1 - \alpha)$$

and differentiating gives  $\alpha^* = \frac{5}{7}$  and

$$V[X_{\alpha^*}] = \frac{21}{49}$$

# Mean and Variance of sample averages

#### Empirical or sample average

Suppose  $Y_1, Y_2, \dots Y_n$  are independent random variables with

$$E[Y_i] = \mu$$
  $V(Y_1) = \sigma^2$ 

Then

$$E\left[\frac{Y_1+Y_2+\cdots Y_n}{n}\right]=\mu$$

and

$$V\left(\frac{Y_1+Y_2+\cdots Y_n}{n}\right)=\frac{\sigma^2}{n}$$

Very important!!