

STAT 315: Normal Random Variables

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Normal Random Variables

PDF of normal random variables

A continuous random variable Y is **normal random variable** with parameters $-\infty < \mu < \infty$ and $\sigma > 0$ if it has the density

$$f(y) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2\sigma^2}} \quad -\infty < y < +\infty$$

We write $Y \sim N(\mu, \sigma^2)$

The CDF

$$F(Y) = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^y e^{-\frac{(y-\mu)^2}{2\sigma^2}}$$

has no closed form. Compute it using technology.

See e.g <https://www.webassign.net/tparise/beta/stats/distributionIndex.html>

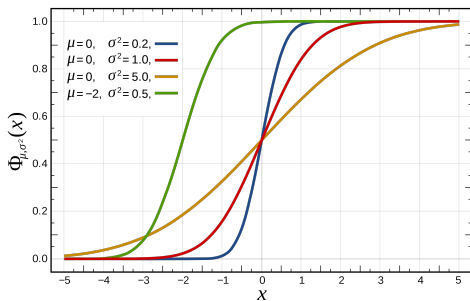
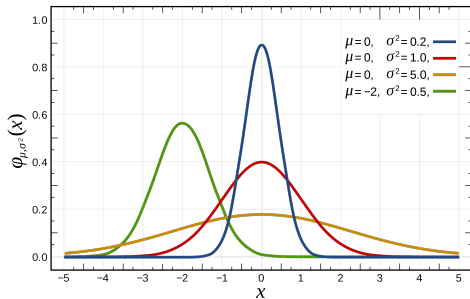
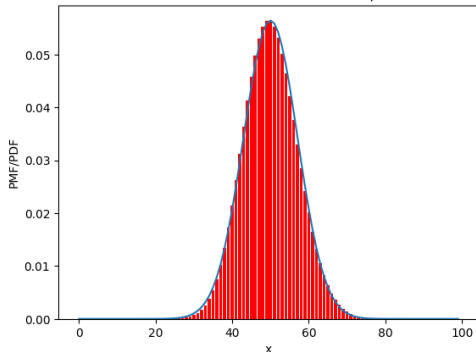


Figure: Top: Pdf, Bottom:Cdf

Why "normal" ?

- Normal random variables are , at least in good approximation.
- Many random variables look very close to a normal (under suitable rescaling), e.g

Poisson PMF ($\lambda = 50$) versus a Normal PDF with $\mu = 50$ and $\sigma^2 = 50$



Y Poisson with large λ

- This come from the Central Limit Theorem (Chapter 7)

Standard normal random variable

Standard normal random variable

A normal random variable is called **standard normal random variable** Z if $\mu = 0$ and $\sigma^2 = 1$. The density is

$$f(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \quad -\infty < y < +\infty$$

Usually we reserve the letter Z is $Z \sim N(0, 1)$

Theorem

If Z is standard normal, $Z \sim N(0, 1)$, then

$$Y = \sigma Z + \mu$$

is normal and $Y \sim N(\mu, \sigma^2)$

Normal-table from your book

848 Appendix 3 Tables

Table 4 Normal Curve Areas
Standard normal probability in right-hand tail
(for negative values of z , areas are found by symmetry)



| z | Second decimal place of z | | | | | | | | | |
|-----|---------------------------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| | .00 | .01 | .02 | .03 | .04 | .05 | .06 | .07 | .08 | .09 |
| 0.0 | .5000 | .4960 | .4920 | .4880 | .4840 | .4801 | .4761 | .4721 | .4681 | .4641 |
| 0.1 | .4602 | .4562 | .4522 | .4483 | .4443 | .4404 | .4364 | .4325 | .4286 | .4247 |
| 0.2 | .4207 | .4168 | .4129 | .4090 | .4052 | .4013 | .3974 | .3936 | .3897 | .3859 |
| 0.3 | .3821 | .3783 | .3745 | .3707 | .3669 | .3632 | .3594 | .3557 | .3520 | .3483 |
| 0.4 | .3446 | .3409 | .3372 | .3336 | .3300 | .3264 | .3228 | .3192 | .3156 | .3121 |
| 0.5 | .3085 | .3050 | .3015 | .2981 | .2946 | .2912 | .2877 | .2843 | .2810 | .2776 |
| 0.6 | .2743 | .2709 | .2676 | .2643 | .2611 | .2578 | .2546 | .2514 | .2483 | .2451 |
| 0.7 | .2420 | .2389 | .2358 | .2327 | .2296 | .2266 | .2236 | .2206 | .2177 | .2148 |
| 0.8 | .2119 | .2090 | .2061 | .2033 | .2005 | .1977 | .1949 | .1922 | .1894 | .1867 |
| 0.9 | .1841 | .1814 | .1788 | .1762 | .1736 | .1711 | .1685 | .1660 | .1635 | .1611 |
| 1.0 | .1587 | .1562 | .1539 | .1515 | .1492 | .1469 | .1446 | .1423 | .1401 | .1379 |
| 1.1 | .1357 | .1335 | .1314 | .1292 | .1271 | .1251 | .1230 | .1210 | .1190 | .1170 |
| 1.2 | .1151 | .1131 | .1112 | .1093 | .1075 | .1056 | .1038 | .1020 | .1003 | .0985 |
| 1.3 | .0968 | .0951 | .0934 | .0918 | .0901 | .0885 | .0869 | .0853 | .0838 | .0823 |
| 1.4 | .0808 | .0793 | .0778 | .0764 | .0749 | .0735 | .0722 | .0708 | .0694 | .0681 |
| 1.5 | .0668 | .0655 | .0643 | .0630 | .0618 | .0606 | .0594 | .0582 | .0571 | .0559 |
| 1.6 | .0548 | .0537 | .0526 | .0516 | .0505 | .0495 | .0485 | .0475 | .0465 | .0455 |
| 1.7 | .0446 | .0436 | .0427 | .0418 | .0409 | .0401 | .0392 | .0384 | .0375 | .0367 |
| 1.8 | .0359 | .0352 | .0344 | .0336 | .0329 | .0322 | .0314 | .0307 | .0301 | .0294 |
| 1.9 | .0287 | .0281 | .0274 | .0268 | .0262 | .0256 | .0250 | .0244 | .0239 | .0233 |
| 2.0 | .0228 | .0222 | .0217 | .0212 | .0207 | .0202 | .0197 | .0192 | .0188 | .0183 |
| 2.1 | .0179 | .0174 | .0170 | .0166 | .0162 | .0158 | .0154 | .0150 | .0146 | .0143 |
| 2.2 | .0139 | .0136 | .0132 | .0129 | .0125 | .0122 | .0119 | .0116 | .0113 | .0110 |
| 2.3 | .0107 | .0104 | .0102 | .0099 | .0096 | .0094 | .0091 | .0089 | .0087 | .0084 |
| 2.4 | .0082 | .0080 | .0078 | .0075 | .0073 | .0071 | .0069 | .0068 | .0066 | .0064 |
| 2.5 | .0062 | .0060 | .0059 | .0057 | .0055 | .0054 | .0052 | .0051 | .0049 | .0048 |
| 2.6 | .0047 | .0045 | .0044 | .0043 | .0041 | .0040 | .0039 | .0038 | .0037 | .0036 |
| 2.7 | .0035 | .0034 | .0033 | .0032 | .0031 | .0030 | .0029 | .0028 | .0027 | .0026 |
| 2.8 | .0026 | .0025 | .0024 | .0023 | .0023 | .0022 | .0021 | .0021 | .0020 | .0019 |
| 2.9 | .0019 | .0018 | .0017 | .0017 | .0016 | .0016 | .0015 | .0015 | .0014 | .0014 |
| 3.0 | .00135 | | | | | | | | | |
| 3.5 | .000233 | | | | | | | | | |
| 4.0 | .0000117 | | | | | | | | | |
| 4.5 | .00000340 | | | | | | | | | |
| 5.0 | .000000287 | | | | | | | | | |

From R. E. Walpole, *Introduction to Statistics* (New York: Macmillan, 1968).

You can always standardize your RV. If Y is normal with mean μ and variance σ then $Z = \frac{Y-\mu}{\sigma}$ is standard normal.

For example

$$\begin{aligned}
 P(a \leq Y \leq b) \\
 &= P\left(\frac{a-\mu}{\sigma} Z \leq \frac{b-\mu}{\sigma}\right)
 \end{aligned}$$

example

- A company that manufactures and bottles apple juice uses a machine that automatically fills 16 ounce bottles. The amount dispensed has been observed to be approximately normally distributed with mean 16 ounces and standard deviation of .5 ounce.
 - ▶ What is the probability that bottles contains less than 14.5 ounces?
 - ▶ Suppose bottle the big enough to contain 17 ounces. What is the probability that the bottles overflow.
 - ▶ How big should the bottles be so less than one percent of bottles overflow?

Mean, variance

Mean and Variance of normal RV

If $Y \sim N(\mu, \sigma^2)$ then

$$E[Y] = \mu \quad V[Y] = \sigma^2$$

Proof: By the previous theorem we can assume $\mu = 0$ and $\sigma^2 = 1$.

$$E[Y] = \int_{-\infty}^{\infty} y \frac{e^{-y^2/2}}{\sqrt{2\pi}} dy = 0 \quad (\text{the integrand is odd})$$

$$\begin{aligned} E[Y^2] &= \int_{-\infty}^{\infty} y \times y \frac{e^{-y^2/2}}{\sqrt{2\pi}} dy = - \int_{-\infty}^{\infty} y \frac{d}{dy} \frac{e^{-y^2/2}}{\sqrt{2\pi}} \\ &\stackrel{\text{IBP}}{=} -y \frac{e^{-y^2/2}}{\sqrt{2\pi}} \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} \frac{e^{-y^2/2}}{\sqrt{2\pi}} = 1 \end{aligned}$$